

Review article

Child sexual abuse and non-suicidal self-injury: meta-analysis

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Background

Many theorists posit that childhood sexual abuse has a central role in the aetiology of self-injurious behaviour. Studies that report statistically significant associations between a history of such abuse and self-injury are cited to support this view.

Aims

A meta-analysis was conducted to determine systematically the magnitude of the association between childhood sexual abuse and self-injurious behaviour.

Method

Forty-five analyses of the association were identified. Effect sizes were converted to a standard metric and aggregated.

Results

The relationship between childhood sexual abuse and self-injurious behaviour is relatively small (mean weighted aggregate $\phi=0.23$). This figure may be inflated owing to publication bias. In studies that statistically controlled for psychiatric risk factors, childhood sexual abuse explained little or no unique variance in self-injurious behaviour.

Conclusions

Theories that childhood sexual abuse has a central or causal role in the development of self-injurious behaviour are not supported by the available empirical evidence. Instead, it appears that the two are modestly related because they are correlated with the same psychiatric risk factors.

Declaration of interest

None.

Self-injurious behaviour can be defined as the causing of intentional, direct damage to one's body tissue without suicidal intent.¹ Common examples include cutting and burning of the skin.^{2–5} Because such behaviour is associated with suicide and psychiatric disorders,^{3,6,7} and because its treatment can be challenging,^{8–11} it has attracted substantial attention in both the clinical and research literature. Although the clinical correlates and functions of self-injurious behaviour have been studied extensively,^{7,12–15} little is known about its aetiology.

Many theorise that childhood sexual abuse has a primary aetiological role. For example, van der Kolk *et al*¹⁶ (p. 1669) wrote that childhood trauma such as sexual abuse 'contributes heavily to the initiation of self-destructive behaviour'. Wonderlich *et al*¹⁷ (p. 203) suggested that individuals subjected to childhood sexual abuse 'engage in a broad array of self-destructive behaviors that may serve to reduce emotional distress associated with their abuse'. Noll *et al*¹⁸ (p. 1467) proposed that sexually abused individuals who self-injure 'may be reenacting the abuse perpetrated on them'. Cavanaugh¹⁹ (pp. 97, 99) described self-injurious behaviour as a 'manifestation of sexual abuse'. Stone²⁰ implicated sexual abuse by a male relative in the development of such behaviour. More recently, Yates²¹ theorised that sexual abuse and other childhood traumas cause emotional and relational vulnerabilities which in turn create the need for self-injurious behaviour as a maladaptive coping strategy.

Those who advocate an aetiological role of childhood sexual abuse point to the numerous studies that document a relationship between histories of such abuse and self-injurious abuse. However, to characterise accurately the empirical relationship between the two variables it is necessary to take into account studies that find small or no associations, in addition to studies that find a positive association. As previous efforts to review the empirical literature on this topic have taken a narrative approach,²² the meta-analysis reported here was conducted to systematically quantify the research findings on the relationship between a history of childhood sexual abuse and the development of self-injurious behaviour.

Method**Inclusion and exclusion criteria**

Studies reporting original research findings regarding the relationship between a history of childhood sexual abuse and self-injurious behaviour were included in this review. Studies in which all participants had histories of childhood sexual abuse or all participants had histories of self-injurious behaviour were excluded, since such studies could not provide measures of association between the two (e.g. Noll *et al*¹⁸). Studies examining self-injurious behaviour with suicidal intent, or that did not distinguish between such behaviour with and without suicidal intent, were also excluded from the meta-analysis (e.g. Romans *et al*,²³ Sansone *et al*,²⁴ Brown *et al*²⁵). Studies examining participants with developmental disabilities or psychosis were excluded. Finally, studies that examined childhood abuse without distinguishing between physical, sexual and other forms of abuse were excluded (e.g. Brodsky *et al*²⁶).

Search strategy

To identify appropriate studies, a literature search was conducted using three database sources: PubMed, PsycINFO, and the Web of Knowledge Science Citation and Social Science Citation Indices. Owing to ambiguity regarding terminology, multiple keywords were identified, and the following search string was used: (self-injury or self-injurious behaviour OR deliberate self-harm OR self-mutilation OR self-mutilative behaviour OR self-destructive) AND (sex abuse OR sexual abuse). Studies published up to the end of June 2006 were surveyed.

Our search strategy yielded 156 empirical English-language studies and these were obtained for further inspection regarding inclusion and exclusion criteria. Of these, 100 were excluded (Fig. 1).

The remaining 56 studies met inclusion criteria. However, for 16 of these there was not enough information to extract an effect size regarding the abuse-behaviour association and efforts to

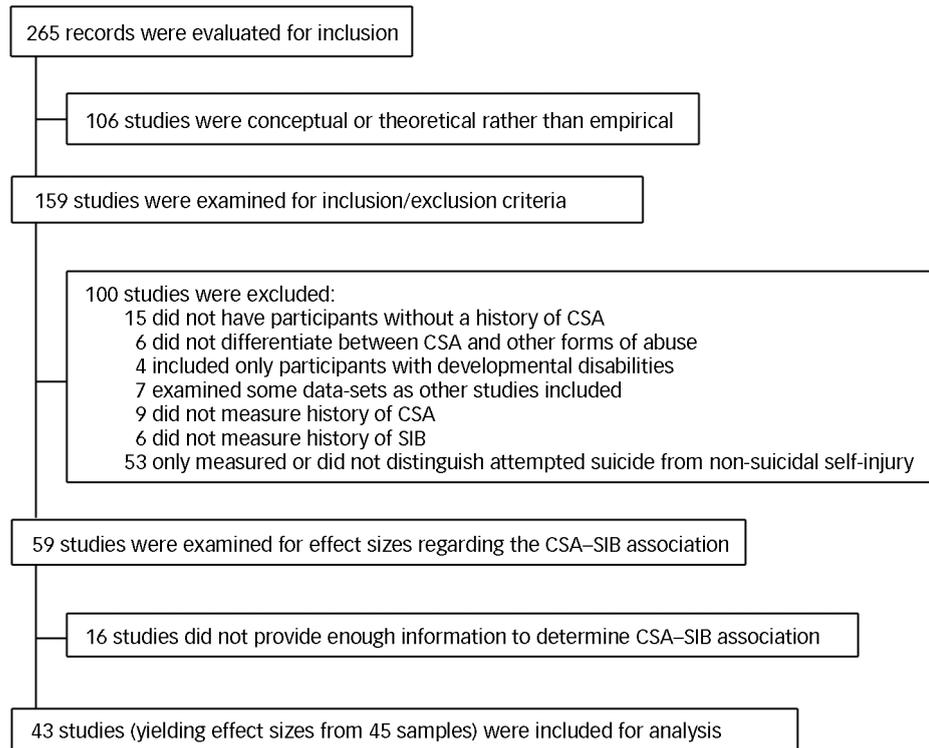


Fig. 1 QUOROM (Quality of Reporting of Meta-analyses) diagram (CSA, childhood sexual abuse; SIB, self-injurious behaviour).

obtain the data from study authors were not successful. The remaining 40 studies with known effect sizes were retained for inclusion in the meta-analysis. We inspected the reference sections of studies meeting inclusion criteria to locate additional relevant studies that might have been missed by our search strategy; only three additional studies meeting inclusion criteria could be located, all of which were published before 1990.^{27–29} Thus we concluded that our search strategy was sufficiently comprehensive and inclusive. Incorporating the three additional studies yielded a total of 43 studies^{5,16,17,27–66} that met full inclusion criteria and were retained for the meta-analysis. These 43 studies contributed effect sizes from 45 independent samples.

Data analysis and study details

For each study in the meta-analysis, effect sizes indicating the relationship between childhood sexual abuse and self-injurious behaviour were extracted or converted to phi coefficient effect sizes. A phi coefficient is a measure of the degree of association between two dichotomous variables and its interpretation is comparable to other correlation coefficients. Methodological details of the 43 studies and 45 samples – including sample size, sample type and demographic variables – are presented in online Table DS1. Meta-analytic analyses were conducted with Comprehensive Meta-Analysis version 2.2.023 (Biostat; Englewood, New Jersey, USA). The effect sizes were examined for heterogeneity and the mean weighted aggregate effect size was computed, adopting a fixed-effects model in the case of a homogeneous distribution of effect sizes and a random-effects model in the case of a heterogeneous distribution. Potential continuous moderators (age, percentage female) of the aggregate effect size were examined with meta-regression and a categorical moderator (type of sample) was examined with an analogue of an analysis of variance procedure appropriate for effect size data. A fixed-effects

model was used when the factors adequately explained the heterogeneity. When additional heterogeneity remained, a mixed-effects model was used.

Results

Table DS1 presents the results from 45 samples regarding the relationship between history of childhood sexual abuse and self-injurious behaviour. The mean weighted aggregate phi coefficient was 0.23 (95% CI 0.20–0.26) using a random effects model, and was significantly different from 0 ($P < 0.001$). Phi coefficients ranged from 0.01 to 0.45 and the distribution exhibited significant heterogeneity ($Q = 90.47, P < 0.001$). Moderator analyses indicated that the magnitude of phi was not related to sample age or gender. Using a mixed-effects model, the type of sample was a significant moderator of the relationship between sexual abuse and self-injurious behaviour ($Q(1, 39) = 5.34, P < 0.05$). This relationship was stronger for the clinical samples ($n = 31; \phi = 0.24$) than for the non-clinical samples ($n = 10; \phi = 0.18$). For this latter analysis, four samples were excluded because they could not be discretely classified as either non-clinical or clinical.^{17,30–32}

We examined the likelihood of publication bias by plotting the standard error as a function of Fisher's Z for each of the 45 effect sizes. On inspection the pattern indicated a lack of symmetry, whereby there were fewer smaller studies with smaller effect sizes in the group located for the review. Kendall's tau was significant ($0.25, P < 0.01$), indicating an association between the treatment effect and the standard error. Similarly, Egger's test of the intercept was significant ($t = 4.82, P < 0.001$). Although these tests are useful for detecting a relationship between sample size and effect size, they cannot isolate the cause, only one of which is publication bias. The fail-safe N indicated that 5462 null studies would need to be located and included to nullify the effect found.

All studies that controlled for psychological risk factors found either minimal or negligible unique associations between childhood sexual abuse and self-injurious behaviour. Because different studies controlled for different variables, these results cannot be statistically aggregated and are thus described here qualitatively. Gratz *et al*⁴⁵ found that the abuse–behaviour relationship became non-significant when controlling for dissociation and several family environment variables (i.e. physical abuse, insecure attachment, emotional neglect and childhood separation), although the relationship remained marginally significant when analyses were limited to female participants. In contrast, Martin *et al*⁵¹ found that the relationship remained statistically significant for male but not female participants when controlling for depression, hopelessness and family functioning. Zoroglu *et al*⁶⁵ found that childhood sexual abuse maintained a statistically significant association with self-injurious behaviour when controlling for dissociation, although the childhood variables neglect, physical abuse and emotional abuse all maintained larger unique associations with such behaviour than did sexual abuse.

All remaining studies that controlled for psychosocial variables found non-significant relationships between childhood sexual abuse and self-injurious behaviour. Evren & Evren³⁹ found that childhood physical abuse (but not sexual abuse) maintained a significant association with self-injurious behaviour when controlling for demographic, family history, and clinical variables. Zlotnick *et al*⁶⁴ found that the abuse–behaviour association was no longer significant when controlling for dissociation, alexithymia and self-destructive behaviours. Zweig-Frank *et al*⁶⁶ reported that the association became non-significant when controlling for family environment variables and a diagnosis of borderline personality disorder. Likewise, in Gladstone *et al*⁴¹ the correlation became non-significant when controlling for borderline personality disorder. Finally, in Parker *et al*⁵⁴ the association became non-significant when controlling for maternal depression, suicidal ideation, current drug use and suicide attempt history.

Discussion

Our meta-analysis examined the association between a history of childhood sexual abuse and the development of self-injurious behaviour. Across 45 samples, the aggregate phi coefficient was 0.23, indicating a relatively small relationship between the two. These results suggest that childhood sexual abuse accounts for no more than 5% of the variance in the development of self-injurious behaviour. Therefore it is unlikely that childhood sexual abuse has a primary role in the development or maintenance of such behaviour.

Significantly, studies with smaller sample sizes tended to report larger relationships. For example, the median phi coefficient for samples with more than 125 participants was 0.21 ($n=22$), compared with a median of $\phi=0.33$ ($n=23$) for samples with 125 or fewer participants. This result suggests the possibility of a bias towards publishing studies with statistically significant results, since studies with smaller sample sizes require larger effect sizes to achieve statistical significance. Indeed, formal analyses found evidence of publication bias, suggesting that smaller studies with positive findings were more likely to be published than smaller studies with null or negative findings.

Finally, childhood sexual abuse appears to explain little or no unique variance in self-injurious behaviour. In studies that controlled for variables such as family environment, dissociation, alexithymia, hopelessness and borderline personality disorder, the abuse–behaviour relationship became minimal or negligible.^{39,45,51,54,64,66} In addition, this relationship was stronger in

clinical samples, in which multiple psychiatric risk factors are likely to be present. Taken as a whole, the pattern of findings suggests that childhood sexual abuse might be best conceptualised as a proxy risk factor for self-injurious behaviour.⁶⁷ In other words, the two might be associated because they are correlated with the same psychiatric risk factors, as opposed to there being a unique or aetiological link between them. At the same time, in some cases childhood sexual abuse might contribute to the initiation of self-injurious behaviour through mediating variables such as depression, anxiety and self-derogation, each of which is known to relate to both childhood sexual abuse and self-injurious behaviour.^{13,68}

Future directions

Variability in the conceptual and operational definitions used by the studies included in the meta-analysis suggests directions for future research. For example, self-injurious behaviour can manifest in many ways and it is possible that the method, frequency, medical severity or other aspects of such behaviour could moderate the abuse–behaviour relationship. Future research should examine this possibility. In addition, meta-analytic data indicate that the association between childhood sexual abuse and psychopathological symptoms tends to be larger for more severe forms of abuse.⁶⁸ Future studies should therefore give consideration to abuse parameters indicative of increased severity (e.g. coercion, frequency, relation to perpetrator, penetration). Initial attempts to examine the relationship of severity parameters to self-injurious behaviour have yielded mixed results.^{22,34,51,63,66} If the most severe forms of childhood sexual abuse are examined, it is possible that the association with self-injurious behaviour might be larger than that reported in this meta-analysis. In the absence of such evidence, however, theories that childhood sexual abuse is a primary cause of such behaviour lack empirical justification.

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Data supplement

Table DS1 Methods and findings of 45 analyses of the relationship between childhood sexual abuse and self-injurious behaviour

Study	Sample size, <i>n</i>	Age, years	Female, %	Sample type	Phi ^a
Bierer <i>et al</i> (2003), Sample 1 ³³	118	38	0	Personality disorder out-patients	0.01
Bierer <i>et al</i> (2003), Sample 2 ³³	64	38	100	Personality disorder out-patients	0.04
Boudewyn & Liem (1995) ³⁴	438	25	61	College students	0.27
Briere & Gil (1998), Sample 1 ³⁵	927	46	50	General population sample	0.13
Briere & Gil (1998), Sample 2 ³⁵	390	36	78	Mixed psychiatric patients	0.25
Briere & Zaidi (1989) ³⁶	50	34	100	Psychiatric emergency room patients	0.24
Brown <i>et al</i> (1999) ³⁷	117	25	98	Eating disorder patients	0.21
Carroll <i>et al</i> (1980) ²⁷	28	28	71	Mixed psychiatric patients	0.35
Craine <i>et al</i> (1988) ²⁸	105	35	100	Mixed psychiatric patients	0.22
Darche (1990) ³⁸	96	15	100	Mixed psychiatric patients	0.32
Evren & Evren (2005) ³⁹	136	36	0	Substance disorder disorder patients	0.27
Favaro & Santonastaso (1999) ⁴⁰	175	24	NA	Bulimia patients	0.24 ^b
Gladstone <i>et al</i> (1999) ⁴¹	171	43	100	Depressed patients	0.27 ^c
Gladstone <i>et al</i> (2004) ⁴²	125	38	100	Depressed patients	0.19 ^c
Gleaves & Eberenz (1993) ⁴³	535	n/a	100	Eating disorder patients	0.20
Gratz (2006) ⁴⁴	200	23	100	College students	0.19
Gratz <i>et al</i> (2002) ⁴⁵	133	23	67	College undergraduates	0.24
Jarvis & Copeland (1997) ⁴⁶	180	33	100	Substance and trauma patients	0.25
Joyce <i>et al</i> (2006) ⁴⁷	195	n/a	57	Depressed patients	0.21 ^d
Kroll <i>et al</i> (1996) ⁴⁸	38	40	100	Mixed psychiatric patients	0.37
Lipschitz <i>et al</i> (1999) ⁴⁹	71	15	52	Mixed psychiatric patients	0.30
Low <i>et al</i> (2000) ⁵⁰	50	32	100	Mixed psychiatric patients	0.37
Martin <i>et al</i> (2004) ⁵¹	2485	14	45	Community sample	0.17
Matsumoto <i>et al</i> (2004) ⁵²	65	24	100	Mixed psychiatric patients	0.32
Nijman <i>et al</i> (1999) ⁵	47	38	48	Mixed psychiatric patients	0.33
Paivio & McCulloch (2004) ⁵³	100	21	100	College students	0.45
Parker <i>et al</i> (2005) ⁵⁴	282	35	74	Depressed patients	0.16 ^e
Pettigrew & Burcham (1997) ⁵⁵	146	33	100	Mixed psychiatric patients	0.11
Rodriguez-Srednicki (2001) ⁵⁶	441	21	100	College students	0.08
Rose <i>et al</i> (1991) ⁵⁷	89	n/a	44	Chronic psychiatric disorder patients	0.44
Sar <i>et al</i> (2004) ⁵⁸	38	33	87	Conversion disorder patients	0.24 ^f
Schwartz <i>et al</i> (1989) ²⁹	60	15	100	Substance disorder patients	0.08 ^g
Swanston <i>et al</i> (1999) ³¹	51	18	91	Abuse victims and controls	0.36
Tobin & Griffing (1996) ⁵⁹	103	29	94	Eating disorder patients	0.42
Tyler <i>et al</i> (2003) ³²	428	17	56	Homeless	0.21
van der Kolk <i>et al</i> (1991) ¹⁶	74	18–39	53	Mixed psychiatric and forensic cases	0.36
Whitlock <i>et al</i> (2006) ⁶⁰	2849	Mostly 18–24	56	Undergraduate and graduate students	0.14 ^h
Wonderlich <i>et al</i> (1996) ¹⁷	65	34	100	Incest victims and psychiatric controls	0.39
Wonderlich <i>et al</i> (2001) ³⁰	51	38	100	Abuse victims and controls	0.33
Wright <i>et al</i> (2004) ⁶¹	524	15	100	Secondary school students	0.26 ⁱ
Ystgaard <i>et al</i> (2004) ⁶²	41	n/a	65	Suicide attempters	0.45
Zanarini <i>et al</i> (2002) ⁶³	290	27	80	Borderline in-patients	0.35
Zlotnick <i>et al</i> (1996) ⁶⁴	148	33	100	Mixed psychiatric patients	0.31
Zoroglu <i>et al</i> (2003) ⁶⁵	818	16	61	High-school students	0.15
Zweig-Frank <i>et al</i> (1994a) ⁶⁶	150	18–48	100	Borderline personality disorder patients	0.17
Mean weighted aggregate					0.23
Unweighted median					0.25

n/a, information not available in the article.

a. Studies used either chi-squared or correlational analyses depending on whether dichotomous or continuous measures of childhood sexual abuse (CSA) and self-injurious behaviour (SIB) were used. In some cases effect sizes were calculated by E.O.K. on the basis of data reported in the original article; the effect sizes from the original studies were then converted into phi coefficients.

b. This effect size indicates the relationship between CSA and what authors termed 'impulsive self-injurious behaviour'.

c. The study separately analysed associations of CSA to current SIB and history of SIB. We used the effect size for history of SIB.

d. Estimated, based on the reported odds ratio of 3.2 for the relationship of CSA and SIB along with limited information in the article about the rates of SIB and CSA in the sample.

e. Parker *et al*⁵⁴ reported data from three samples of patients with depression. We treated the three samples as one large sample for the purpose of this meta-analysis. In addition, two different effect sizes were calculated for two different forms of CSA: that perpetrated by a parent and that perpetrated by someone other than a parent. These two forms of CSA corresponded to phi coefficients of 0.18 and 0.14 respectively; the average of these two coefficients, 0.16, was used for the purposes of the meta-analysis.

f. Data necessary to calculate the effect size were not reported in the original article and were obtained from the corresponding author by email.

g. Two different effect sizes were calculated for two different forms of CSA: incest and rape corresponded to phi coefficients of -0.06 and 0.22 respectively. The average of these two coefficients, 0.08,⁵⁹ was used for the purposes of the meta-analysis.

h. Whitlock *et al*⁶⁰ divided participants who self-injured into more than one category for most of their analyses. The effect size reported here represents the relationship between CSA and SIB where both are treated as dichotomous, present-absent variables.

i. Based on data from girls in a secondary school setting with and without a history of CSA. A third sample consisting of sexually abused girls from a clinical setting was not included since there was not a sample of non-abused girls from a clinical setting to serve as a comparison group.